

# CONVERGENCE RATES IN THE CENTRAL LIMIT THEOREM FOR THE SUMS OF A RANDOM NUMBER OF INDEPENDENT AND IDENTICALLY DISTRIBUTED RANDOM VARIABLES

Tomoichi NAKATA

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## 1. Introduction and Results.

Let  $\{X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables with  $EX_1=0$ ,  $EX_1^2=1$ , and distribution function  $F$ . Let  $\{N_n, n \geq 1\}$  be a sequence of positive integer-valued random variables independent of  $\{X_n, n \geq 1\}$ . The distribution function of a random variable  $N_n$  is determined by the values  $p_k(n)=P(N_n=k)$   $k \geq 1$ , where  $p_k(n)$  are functions of  $n$  such that for all  $n$ ,  $p_k(n) \geq 0$  and  $\sum_{k=1}^{\infty} p_k(n)=1$ . Put

$$a_n = EN_n, a_0 = 0, M_n^2 = \sigma^2 N_n \text{ (assumed to be finite for all } n).$$

We shall assume that  $a_0 < a_1 < a_2 \dots$ , and  $\lim_{n \rightarrow \infty} a_n = \infty$ .

Further on let

$$S_n = \sum_{j=1}^n X_j, S_{N_n} = \sum_{j=1}^{N_n} X_j, G(x) = P(|X_1| \leq x), F_n(x) = P(S_{N_n} < x \sqrt{N_n}),$$

and denote by  $\Phi(x)$  the standard normal distribution. By  $C$  we shall denote positive constants which are in general different.

The purpose of this paper is to give an extension of Rychlik and Szynal [9] for the sums of a random number of independent and identically distributed random variables.

The obtained results are extensions of some theorems given in [4] [5] and [9].

We shall show the following theorems.

**Theorem 1.** *The following inequality holds,*

$$\sup_x |F_n(x) - \Phi(x)| \leq C[a_n^{-1/2} \int_{|u| \leq \sqrt{a_n}} |u|^3 dF(u) + \int_{|u| > \sqrt{a_n}} u^2 dF(u) + M_n a_n^{-2}]$$

**Theorem 2.** For all  $x$ , the following inequality holds,

$$|F_n(x) - \Phi(x)| \leq C[(1 + |x|)^{-3} \int_0^{(1+|x|)\sqrt{a_n}} \int_{|u|>v} u^2 dF(u) dv + (1 + |x|)^{-2} M_n a_n^{-2}]$$

**Theorem 3.** For all  $x$ , the following inequality holds,

$$|F_n(x) - \Phi(x)| \leq C(1 + |x|^2)^{-1} [a_n^{-1/2} \int_0^{\sqrt{a_n}} u^3 dG(u) + \int_{\sqrt{a_n}}^{\infty} u^2 dG(u) + M_n a_n^{-2}].$$

Heyde [5] has shown the inequality in the case of non random number of Theorem 3. Some applications of Theorem 3 are given in following corollaries which extend Theorem 4 and Theorem 5 of Rychlik and Szynal [9]. They have shown the case of  $p = \infty$  in following corollaries.

**Corollary 1.** Let  $\{X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables such that  $EX_1 = 0$ ,  $EX_1^2 = 1$  and  $E|X_1|^{2+\delta} < \infty$ , where  $0 < \delta < 1$ . If there exist positive constants  $C_0$  and  $C_1$  such that for every  $n \geq 2$

$$C_0 \leq a_n \leq C_1 a_{n-1} \quad (1)$$

and if

$$\sum_{n=1}^{\infty} M_n a_n^{-2+\delta/2} < \infty, \quad \sum_{k=n}^{\infty} a_k^{(\delta-3)/2} = o(a_n^{(\delta-1)/2}),$$

then we have

$$\sum_{n=1}^{\infty} a_n^{-1+\delta/2} \|F_n(x) - \Phi(x)\|_p < \infty \quad (1 \leq p \leq \infty)$$

where  $\|F_n(x) - \Phi(x)\|_p = (\int |F_n(x) - \Phi(x)|^p dx)^{1/p} \quad (1 \leq p \leq \infty)$

$$\|F_n(x) - \Phi(x)\|_{\infty} = \sup_x |F_n(x) - \Phi(x)|.$$

**Corollary 2.** Let  $\{X_n, n \geq 1\}$  be a sequence of independent and identically distributed random variables such that  $EX_1 = 0$ ,  $EX_1^2 = 1$ ,  $EX_1^2 \log X_1 < \infty$ . Suppose that (1) and if

$$\sum_{n=1}^{\infty} M_n a_n^{-2} < \infty, \quad \sum_{k=n}^{\infty} a_k^{-3/2} = o(a_n^{-1/2}),$$

then we have

$$\sum_{n=1}^{\infty} a_n^{-1} \|F_n(x) - \Phi(x)\|_p < \infty, \quad (1 \leq p \leq \infty).$$

## 2. Proofs of Theorem 1 and Theorem 2.

Now we shall prove Theorem 1.

We have the following inequality from (10) of Rychlik and Szynal [9] p.686 ;

$$\sup |F_n(x) - \Phi(x/A_n)| \leq C[a_n P(|X_1| \geq \sqrt{a_n})$$

$$+ \sqrt{a_n} \int_{|u| \leq \sqrt{a_n}} u dF(u) + a_n^{-1} \int_{|u| > \sqrt{a_n}} |u|^3 dF(u) + M_n a_n^{-2}], \quad (2)$$

where  $A_n^2 = \{ \int_{|u| < \sqrt{a_n}} u^2 dF(u) - (\int_{|u| < \sqrt{a_n}} u dF(u))^2 \}$ .

It is easily seen that the first and second terms in the right hand side of the inequality (2) are dominated by

$$\int_{|u| \leq \sqrt{a_n}} u^2 dF(u). \quad (3)$$

In order to estimate  $|\Phi(x/A_n) - \Phi(x)|$ , we use the following result of Erickson [3] p.527 ;

$$|\Phi(ax) - \Phi(x)| \leq 4/5 \cdot |1 - a^{-1}| \quad (a > 0).$$

Therefore we have

$$|\Phi(x/A_n) - \Phi(x)| \leq |1 - A_n^2| \leq 2 \int_{|u| \leq \sqrt{a_n}} u^2 dF(u). \quad (4)$$

From (2) (3) and (4), we have the inequality of the Theorem 1. Thus the proof of Theorem 1 is complete.

Now we prove Theorem 2.

We define

$$X_{nj} = X_{nj}(x) = \begin{cases} X_j & \text{if } |X_j| \leq (1 + |x|)\sqrt{a_n} \\ 0 & \text{otherwise} \end{cases}$$

for  $n, j=1, 2, 3, \dots$ .

Let us put

$$S_{nk} = \sum_{j=1}^k X_{nj}, \quad \mu_{nk} = kEX_{n1}, \quad B_{nk}^2 = k\sigma^2 X_{n1}.$$

Further we define  $K_i$  ( $i=1, 2, 3, 4$ ) in the following way ;

$$K_1 = \{k; |\mu_{nk}/\sqrt{k}| \geq |x|/2, \quad k > a_n/2\},$$

$$K_2 = \{k; |\mu_{nk}/\sqrt{k}| < |x|/2, \quad k > a_n/2\},$$

$$K_3 = \{k; |\mu_{nk}/\sqrt{k}| \leq |x|/2, k \leq a_n/2\},$$

$$K_4 = \{k; |\mu_{nk}/\sqrt{k}| < |x|/2, k \leq a_n/2\}.$$

For  $|x| < 1$ , the required inequality follows from the theorem 1. Therefore it is sufficient to limit our considerations to the case  $|x| \geq 1$ , we have the inequality for all  $x$ .

$$\begin{aligned} |F_n(x) - \Phi(x)| &\leq \sum_{k=1}^{\infty} p_k(n) |P(S_k < x\sqrt{k}) - \Phi(x)| \\ &\leq \sum_{k \in K_1} p_k(n) |\cdot| + \sum_{k \in K_2} p_k(n) |\cdot| + \sum_{k \in K_3} p_k(n) |\cdot| + \\ &\quad + \sum_{k \in K_4} p_k(n) |\cdot|, \\ &= I_1 + I_2 + I_3 + I_4, \text{ say,} \end{aligned} \quad (5)$$

where  $|\cdot| = |P(S_k < x\sqrt{k}) - \Phi(x)|$ .

Now we treat  $I_1$ . Since  $|\cdot| \leq 1$ , we have

$$\begin{aligned} I_1 &= \sum_{k \in K_1} p_k(n) \leq 2 \sum_{k \in K_1} p_k(n) |\mu_{nk}/x\sqrt{k}| \leq 2\sqrt{2}x^{-1}a_n^{-1/2}E|\mu_{nu}|, \\ &\leq 4\sqrt{2}(1+|x|)^{-1}a_n^{-1/2}a_n \int_{|u| > (1+|x|)\sqrt{a_n}} u dF(u), \\ &\leq 4\sqrt{2}(1+|x|)^{-2} \int_{|u| > (1+|x|)\sqrt{a_n}} u^2 dF(u). \end{aligned} \quad (6)$$

Here we have used the property of  $K_1$ .

To estimate  $I_2$ , we need the following inequality for all  $x$

$$\begin{aligned} I_2 &\leq \sum_{k \in K_2} p_k(n) |P(S_k < x\sqrt{k}) - P(S_{nk} < x\sqrt{k})| + \\ &\quad + \sum_{k \in K_2} p_k(n) |P\left(\frac{S_{nk} - \mu_{nk}}{B_{nk}} < \frac{x\sqrt{k} - \mu_{nk}}{B_{nk}}\right) - \Phi\left(\frac{x\sqrt{k} - \mu_{nk}}{B_{nk}}\right)| + \\ &\quad + \sum_{k \in K_2} p_k(n) |\Phi\left(\frac{x\sqrt{k} - \mu_{nk}}{B_{nk}}\right) - \Phi\left(\frac{x\sqrt{k}}{B_{nk}}\right)| + \\ &\quad + \sum_{k \in K_2} p_k(n) |\Phi\left(\frac{x\sqrt{k}}{B_{nk}}\right) - \Phi(x)|, \\ &= J_1 + J_2 + J_3 + J_4, \text{ say.} \end{aligned} \quad (7)$$

Since  $P(S_k < x\sqrt{k}) \leq P(S_{nk} < x\sqrt{k}) + P(S_k \neq S_{nk})$ , we observe that

$$\begin{aligned} J_1 &= \sum_{k=1}^{\infty} p_k(n) P(S_k \neq S_{nk}) \leq a_n P(|X_1| > (1+|x|)\sqrt{a_n}) \\ &\leq a_n \int_{|u| > (1+|x|)\sqrt{a_n}} dF(u), \\ &\leq (1+|x|)^{-2} \int_{|u| > (1+|x|)\sqrt{a_n}} u^2 dF(u). \end{aligned} \quad (8)$$

To estimate  $J_2$ , using the inequality of Nagaev [6];

$$|P(S_k < x\sqrt{k}) - \Phi(x)| \leq C(1 + |x|^3)^{-1} \sqrt{k} |E| X_1|^3,$$

we have

$$J_2 \leq 4C \sum_{k \in K_2} p_k(n) (1 + |y_{nk}|^3)^{-1} B_{nk}^{-3} k |E| X_1|^3,$$

whence  $y_{nk} = (x\sqrt{k} - \mu_{nk})B_{nk}^{-1}$ . However since

$$(1 + |y_{nk}|^3)^{-1} \leq B_{nk}^3 k^{-3/2} |x - \mu_{nk} k^{-1/2}|^{-3} \leq 8B_{nk}^3 |x|^{-3} k^{-3/2},$$

we have

$$\begin{aligned} J_2 &\leq 32C \sum_{k \in K_2} p_k(n) |x|^{-3} k |E| X_{n1}|^3 a_n^{-3/2}, \\ &\leq 64C \sum_{k \in K_2} p_k(n) |x|^{-3} k a_n^{-3/2} |E| X_{n1}|^3, \\ &\leq 64\sqrt{2} C |x|^{-3} a_n^{-1/2} |E| X_{n1}|^3. \end{aligned} \quad (9)$$

Here we have used the property of  $K_2$ .

As to  $J_3$ , since  $\exp(u^2/2) > u^2/2$  for  $u > 0$ , we have

$$\begin{aligned} J_3 &\leq \sum_{k \in K_2} p_k(n) 2/\sqrt{2\pi} \cdot \left| \int_{(x\sqrt{k} - \mu_{nk})B_{nk}^{-1}}^{x\sqrt{k} B_{nk}^{-1}} u^{-2} du \right|, \\ &\leq \sqrt{2}/\sqrt{\pi} \cdot \sum_{k \in K_2} p_k(n) | -B_{nk}(x\sqrt{k})^{-1} + B_{nk}(x\sqrt{k} - \mu_{nk})^{-1} |, \\ &\leq 4/\sqrt{2\pi} \cdot \sum_{k \in K_2} p_k(n) x^{-2} k^{-1/2} |\mu_{nk}|, \\ &\leq \sqrt{2}/\sqrt{\pi} \cdot |x|^{-1} a_n^{-1/2} \sum_{k \in K_2} p_k(n) k |E X_{n1}|, \\ &\leq 2\sqrt{2}/\sqrt{\pi} \cdot (1 + |x|)^{-2} \int_{|u| > (1 + |x|)\sqrt{a_n}} u^2 dF(u). \end{aligned} \quad (10)$$

Here we have used the property of  $K_2$ .

Since  $\exp(u^2/2) > u^3/8$ , we have

$$\begin{aligned} J_4 &\leq 8/\sqrt{2\pi} \cdot \sum_{k \in K_2} p_k(n) \left| \int_x^{x\sqrt{k} B_{nk}^{-1}} u^{-3} du \right|, \\ &\leq \sqrt{2}/\sqrt{\pi} \cdot \sum_{k \in K_2} p_k(n) x^{-2} (k - B_{nk}^2) k^{-1} \\ &\leq 4\sqrt{2}/\sqrt{\pi} \cdot (1 + |x|)^{-2} \int_{|u| > (1 + |x|)\sqrt{a_n}} u^2 dF(u). \end{aligned} \quad (11)$$

Hence from (7) (8) (9) (10) and (11), choosing a suitable constant, we have

$$\begin{aligned} I_2 &\leq C(1 + |x|)^{-2} \int_{|u| > (1 + |x|)\sqrt{a_n}} u^2 dF(u) + \\ &\quad + C(1 + |x|)^{-3} a_n^{-1/2} \int_{|u| \leq (1 + |x|)\sqrt{a_n}} |u|^3 dF(u). \end{aligned} \quad (12)$$

Now we treat  $I_3$ . Since  $|\cdot| \leq 1$ , we have

$$I_3 \leq 2 \sum_{k \in K_3} p_k(n) |\mu_{nk}/x\sqrt{k}|.$$

However

$$|EX_{n1}| = \left| \int_{|u| \geq (1+|x|)\sqrt{a_n}} u dF(u) \right| \leq (1+|x|)^{-1} a_n^{-1/2},$$

hence

$$|\mu_{nk}| \leq (1+|x|)^{-1} k a_n^{-1/2} \leq |x|^{-1} a_n^{-1/2} k.$$

Therefore we have

$$\begin{aligned} I_3 &\leq 2 \sum_{k \in K_3} p_k(n) |x|^{-2} a_n^{-1/2} k^{1/2} \leq \sqrt{2} |x|^{-2} \sum_{k \in K_3} p_k(n), \\ &\leq 2 \sqrt{2} (1+|x|)^{-2} \sum_{k; k \leq a_n/2} p_k(n), \\ &\leq 4 \sqrt{2} (1+|x|)^{-2} M_n a_n^{-2}. \end{aligned} \quad (13)$$

At last we treat  $I_4$ . We need the inequality (7) replacing  $K_2$  by  $K_4$ .

$$I_4 \leq J_1 + J_2 + J_3, \quad \text{say,} \quad (14)$$

where

$$J_3 = \sum_{k \in K_4} p_k(n) |\Phi((x\sqrt{k} - \mu_{nk})B_{nk}^{-1}) - \Phi(x)|.$$

From (12), we have

$$\begin{aligned} J_1 &= \sum_{k \in K_4} p_k(n) |P(S_k < x\sqrt{k}) - P(S_{nk} < x\sqrt{k})|, \\ &\leq (1+|x|)^{-2} \int_{|u| > (1+|x|)\sqrt{a_n}} u^2 dF(u). \end{aligned} \quad (15)$$

In order to estimate  $J_2$ , we use the following result of Agnew ([1] Theorem 2.2) that if  $F(x)$  and  $G(x)$  and two distribution functions with zero mean and unit variance then

$$|F(x) - G(x)| \leq (1+x^2)^{-1} \quad -\infty < x < \infty.$$

Then we have

$$J_2 \leq \sum_{k \in K_4} p_k(n) (1 + |y_{nk}|^2)^{-1},$$

where  $y_{nk} = (x\sqrt{k} - \mu_{nk})B_{nk}^{-1}$ .

However since

$$(1 + |y_{nk}|^2)^{-1} \leq B_{nk}^2 k^{-1} (x - \mu_{nk} k^{-1/2})^{-2} \leq 4 B_{nk}^2 k^{-1} x^{-2} + 4 x^{-2},$$

we have

$$\begin{aligned} J_2 &\leq 8 (1+|x|)^{-2} \sum_{k; k \leq a_n/2} p_k(n) \\ &\leq 16 (1+|x|)^{-2} M_n a_n^{-2}. \end{aligned} \quad (16)$$

As to  $J_3$ , since  $\exp(u^2/2) > u^3/8$ , for  $u > 1$ , we have

$$\begin{aligned} J_3 &\leq 8 \sum_{k \in K_4} p_k(n) \left| \int_x^{(x\sqrt{k} - \mu_{nk})B_{nk}^{-1}} B_{nk}^{-1} u^{-3} du \right|, \\ &\leq 8 \sum_{k \in K_4} p_k(n) | -B_{nk}^2 (x\sqrt{k} - \mu_{nk})^{-2} 2^{-1} + (2x^2)^{-1} |, \end{aligned}$$

$$\begin{aligned} &\leq 24(1+|x|)^{-2} \sum_{k; k \leq a_n/2} p_k(n), \\ &\leq 48(1+|x|)^{-2} M_n a_n^{-2}. \end{aligned} \quad (17)$$

Hence from (14) (15) (16) and (17), we have

$$I_4 \leq C(1+|x|)^{-2} \int_{|u| > (1+|x|)\sqrt{a_n}} u^2 dF(u) + C(1+|x|)^{-2} M_n a_n^{-2}. \quad (20)$$

From (5) (6) (12) (13) and (20), choosing a suitable constant, we have the following inequality

$$\begin{aligned} |F_n(x) - \Phi(x)| &\leq C[(1+|x|)^{-3} a_n^{-1/2} \int_{|u| \leq (1+|x|)\sqrt{a_n}} |u|^3 dF(u) + \\ &\quad + (1+|x|)^{-2} \int_{|u| > (1+|x|)\sqrt{a_n}} u^2 dF(u) + (1+|x|)^{-2} M_n a_n^{-2}] \\ &=: D_1 + D_2 + D_3, \text{ say.} \end{aligned} \quad (21)$$

Being noted that

$$\begin{aligned} \int_0^{(1+|x|)\sqrt{a_n}} \int_{|u| > v} u^2 dF(u) dv &= \int_{|u| < (1+|x|)\sqrt{a_n}} u^3 dF(u) + \\ &\quad + (1+|x|)\sqrt{a_n} \int_{|u| > (1+|x|)\sqrt{a_n}} u^2 dF(u), \end{aligned}$$

we have the inequality in Theorem 2.

Thus the proof of Theorem 2 is complete.

### 3. Proofs of Theorem 3 and Corollaries.

In order to prove Theorem 3, we use Theorem 2. In the inequality (21), we have

$$\begin{aligned} D_1 &= C(1+|x|)^{-3} a_n^{-1/2} \int_0^{(1+|x|)\sqrt{a_n}} u^3 dG(u), \\ &\leq C(1+|x|)^{-3} a_n^{-1/2} \int_0^{\sqrt{a_n}} u^3 dG(u) + \\ &\quad + C(1+|x|)^{-2} \int_{\sqrt{a_n}}^{(1+|x|)\sqrt{a_n}} u^2 dG(u). \end{aligned}$$

$$\text{Since } D_2 \leq C(1+|x|)^{-2} \int_{(1+|x|)\sqrt{a_n}}^{\infty} u^2 dG(u),$$

we have, noting  $1+|x| \leq (1+|x|)^2$

$$\begin{aligned} D_1 + D_2 &\leq C(1+|x|)^{-3} a_n^{-1/2} \int_0^{\sqrt{a_n}} u^3 dG(u) + \\ &\quad + C(1+|x|)^{-2} \int_{\sqrt{a_n}}^{\infty} u^2 dG(u). \end{aligned}$$

Therefore we have

$$D_1 + D_2 + D_3 \leq C(1 + |x|^2)^{-1} [a_n^{-1/2} \int_0^\infty a_n u^3 dG(u) + \int_0^\infty \sqrt{a_n} u^2 dG(u) + M_n a_n^{-2}].$$

Thus we complete the proof of Theorem 3.

To show Corollary 1 and Corollary 2, we first observe that the series

$$\sum_{n=1}^{\infty} a_n^{-1+\delta/2} \sup_x |F_n(x) - \Phi(x)|, \quad \sum_{n=1}^{\infty} a_n^{-1} \sup_x |F_n(x) - \Phi(x)|$$

converge. However the convergences of these series are given by Theorem 2 and Theorem 3 of Rychlik and Szynal [9] in view of Theorem 1. For  $1 \leq p < \infty$ , from Theorem 3, we have

$$\|F_n(x) - \Phi(x)\|_p \leq C_n (\int |1 + |x|^2|^{-p} dx)^{1/p}$$

where

$$C_n = C[a_n^{-1/2} \int_0^\infty a_n u^3 dG(u) + \int_0^\infty \sqrt{a_n} u^2 dG(u) + M_n a_n^{-2}].$$

Hence the convergences of series

$$\sum_{n=1}^{\infty} a_n^{-1+\delta/2} \|F_n(x) - \Phi(x)\|_p, \quad \sum_{n=1}^{\infty} a_n^{-1} \|F_n(x) - \Phi(x)\|_p$$

are given by the proof of the case  $p = \infty$ .

## REFERENCES

- [1] Agnew, R. P. (1957). Estimates for global central limit theorems. Ann. Math. Stat. 28, 26-42.
- [2] Bikyalis, A. (1966). Estimates for the remainder term in the central limit theorem. Litovsk. Matem. Sb. 6, 321-346.
- [3] Erickson, R. V. (1974).  $L_1$  bounds for asymptotic normality of  $m$ -dependent sums using Stein's technique. Ann. Prob. 2, 522-529.
- [4] Galestyan, F. N. (1971). On the rate of convergence in the central limit theorem. Theo. Prob. Math. Stat. 18, 175-180.
- [5] Heyde, C. C. (1975). A nonuniform bound on convergence to normality. Ann. Prob. 3, 5, 903-907.
- [6] Nagaev, S. V. (1976). Some limit theorems for large deviations. Theo. Prob. Appl. 10, 214-235.
- [7] Nakata, T. (1976). On the rate of convergence in mean central



limit theorem for martingale differences. Rep. Stat. Appl. Res. JUSE, 24, 3.

- [8] Nakata, T. (1977) A nonuniform bound on convergenc to normality for independent random variables. Bull. Fac. Lib. Arts. Chukyo Univ. 18, 19-24.
- [9] Rychlik, Z. and Szynal, D. (1974). On the convergence rates in the central limit theorem for the sums of a random number of i. i. d. r. v's. Bull. Acad. Poln. Sci., Ser. Sci. Math. Astronom. Phys., 22. 7, 683-690.
- [10] Rychlik, Z. and Szynal, D. (1975). Convergence rates in the central limit theorem for sums of a random number of independent r. v's. Theo. Prob. Appl. 20, 351-362.

Facilty of Liberal Arts  
Chukyo University  
Yagoto Showa-ku  
Nagoya, Japan.